



Ch8: Bivariate Distributions

8.1 Joint distributions of two random variables

◆ We now consider two or more random variables that are defined simultaneously on the same sample space.

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Functions
Functions

Joint Probability Mass Functions

◆ Definition

- *Let X and Y be two discrete random variables defined on the same sample space.*
- *Let the sets of possible values of X and Y be A and B , respectively. The function*

$$p(x, y) =$$

*is called the **joint probability mass function** of X and Y .*

- ◆ Note that $p(x, y) \geq 0$. If $x \notin A$ or $y \notin B$, then
Also,

◆ Definition

- *Let X and Y have joint probability mass function $p(x, y)$.*
- *Let A be the set of possible values of X and B be the set of possible values of Y .*
- *Then the functions $p_X(x) =$ and $p_Y(y) =$ are called, respectively, the **functions of X and Y .***

Example 8.1

- ◆ A small college has 90 male and 30 female professors. An ad hoc committee of five is selected at random to write the vision and mission of the college.
- ◆ Let X and Y be the number of men and women on this committee, respectively.
 - (a)** Find the joint probability mass function of X and Y .
 - (b)** Find p_X and p_Y , the marginal probability mass functions of X and Y .

Example 8.1

(a)

1. The set of possible values for both X and Y is $\{0, 1, 2, 3, 4, 5\}$.

2.

$$p(x, y) = \begin{cases} 1 & \text{if } x, y \in \{0, 1, 2, 3, 4, 5\}, x + y = 5 \\ 0 & \text{otherwise} \end{cases}$$

(b)

1.

Since $p(x, y) = 0$ if $x + y \neq 5$, $\sum_{y=0}^5 p(x, y) =$

and $\sum_{x=0}^5 p(x, y) =$

$p_X(x) =$, $p_Y(y) =$ $x, y \in \{0, 1, 2, 3, 4, 5\}$

- ◆ Let X and Y be discrete random variables with joint probability mass function $p(x, y)$.
- ◆ Let the sets of possible values of X and Y be A and B , respectively.
- ◆ To find $E(X)$ and $E(Y)$, first we calculate p_X and p_Y , the marginal probability mass functions of X and Y , respectively. Then we will use the following formulas.

$$E(X) = \quad ; \quad E(Y) =$$

Example 8.3

◆ Let the joint probability mass function of random variables X and Y be given by

$$p(x, y) = \begin{cases} \frac{1}{70} x(x + y) & \text{if } x = 1, 2, 3, \quad y = 3, 4 \\ 0 & \text{elsewhere} \end{cases}$$

Find $E(X)$ and $E(Y)$.

1. First we need to calculate

2.

$$px(x) = \frac{1}{70} x(x + 3) + \frac{1}{70} x(x + 4) = \frac{1}{35} x^2 + \frac{1}{10} x$$

$$py(y) = \frac{1}{5} + \frac{3}{35} y$$

Example 8.3

3.

$$E(X) = \sum_{x=1}^3 x \left(\frac{1}{35} x^2 + \frac{1}{10} x \right) = \frac{17}{7} \approx 2.43$$

$$E(Y) = \sum_{y=3}^4 y \left(\frac{1}{5} + \frac{3}{35} y \right) = \frac{124}{35} \approx 3.54$$

Theorem 8.1

- ◆ *Let $p(x, y)$ be the joint probability mass function of discrete random variables X and Y .*
- ◆ *Let A and B be the set of possible values of X and Y , respectively.*
- ◆ *If h is a function of two variables from \mathbf{R}^2 to \mathbf{R} , then $h(X, Y)$ is a discrete random variable with the expected value given by*

$$E[h(X, Y)] =$$

provided that the sum is absolutely convergent.

(generalization of Theorem 4.2)

Corollary

◆ *For discrete random variables X and Y ,*
 $E(X + Y) = E(X) + E(Y)$.

Proof: In Theorem 8.1 let $h(x,y)=x+y$. Then

$$\begin{aligned} E(X + Y) &= \\ &= \\ &= E(X) + E(Y) \end{aligned}$$

Joint Probability Density Functions

◆ Definition

Two random variables X and Y , defined on the same sample space, have a continuous joint distribution if there exists a nonnegative function of two variables, $f(x, y)$ on $\mathbf{R} \times \mathbf{R}$, such that for any region R in the xy -plane that can be formed from rectangles by a countable number of set operations,

$$P((X, Y) \in R) =$$

*The function $f(x, y)$ is called the **joint probability density function** of X and Y .*

◆ Let $R = \{(x, y) : x \in A, y \in B\}$, where A and B are *any* subsets of real numbers that can be constructed from intervals by a countable number of set operations. Then (8.2) gives

$$P(X \in A, Y \in B) = \int_B \int_A f(x, y) dx dy$$

Letting $A = (-\infty, \infty)$, $B = (-\infty, \infty)$, (8.3) implies the relation

Definition

◆ Let X and Y have joint probability density function $f(x, y)$; then the functions

$$f_X(x) =$$

$$f_Y(y) =$$

*are called, respectively, the **marginal probability density functions** of X and Y .*

- ◆ Let X and Y be two random variables (discrete, continuous, or mixed).
- ◆ The **joint probability distribution function**, or *joint cumulative probability distribution function*, or simply the *joint distribution of X and Y* , is defined by

$$F(t,u) =$$

for all $-\infty < t, u < \infty$.

- ◆ The **marginal probability distribution function of X , F_X** , can be found from F as follows:

$$F_X(t) = \equiv F(t, \infty).$$

- ◆ Similarly, F_Y , the **marginal probability distribution function of Y** , is

$$F_Y(u) = P(Y \leq u) \equiv$$

Suppose that the joint density function of X and Y is $f(x, y)$

$$F(x, y) = P(X \leq x, Y \leq y) =$$

Assuming that the partial derivatives of F exist, we get

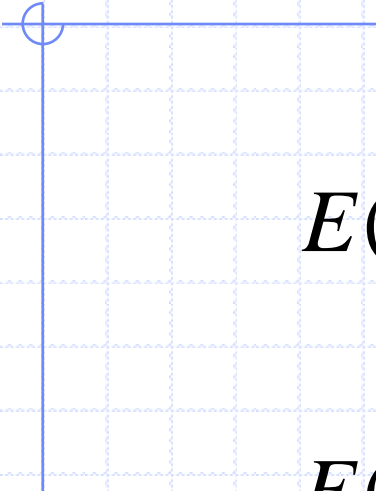
$$f(x, y) =$$

Moreover, $F_X(x) =$

$$= \int_{-\infty}^x \left(\int_{-\infty}^{\infty} f(t, u) du \right) dt = \int_{-\infty}^x f_X(t) dt,$$

and similarly, $F_Y(y) =$

$$\text{And } F'_X(x) = f_X(x), \quad F'_Y(y) = f_Y(y)$$


$$E(X) =$$

$$E(Y) =$$

Example 8.4

- ◆ The joint probability density function of random variables X and Y is given by

$$f(x, y) = \begin{cases} \lambda xy^2 & 0 \leq x \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

- (a)** Determine the value of λ .
(b) Find the marginal probability density functions of X and Y .
(c) Calculate $E(X)$ and $E(Y)$.

Example 8.4

(a) To find λ , note that

Therefore,

$$\begin{aligned} 1 &= \int_0^1 \left(\int_x^1 y^2 dy \right) \lambda x dx = \int_0^1 \left[\frac{1}{3} y^3 \right]_x^1 \lambda x dx = \lambda \int_0^1 \left(\frac{1}{3} - \frac{1}{3} x^3 \right) x dx \\ &= \frac{\lambda}{3} \int_0^1 (1 - x^3) x dx = \frac{\lambda}{3} \int_0^1 (x - x^4) dx = \frac{\lambda}{3} \left[\frac{1}{2} x^2 - \frac{1}{5} x^5 \right]_0^1 = \frac{\lambda}{10} \end{aligned}$$

hence $\lambda = 10$.

$$(b) \quad f_X(x) = \left[\frac{10}{3} x y^3 \right]_x^1 = \frac{10}{3} x (1 - x^3)$$

$$f_Y(y) = \left[5 x^2 y^2 \right]_0^y = 5 y^4$$

$$(c) \quad E(X) = \frac{5}{9}; \quad E(Y) = \int_0^1 y \cdot 5 y^4 dy = \frac{5}{6}$$

Example 8.5

◆ For $\lambda > 0$, let

$$F(x, y) = \begin{cases} 1 - \lambda e^{-\lambda(x+y)} & \text{if } x > 0, y > 0 \\ 0 & \text{otherwise.} \end{cases}$$

Determine if F is the joint probability distribution function of two random variables X and Y .

If F is the joint density of X and Y then

is the joint density function of X and Y . But

$$\frac{\partial^2}{\partial x \partial y} F(x, y) = \begin{cases} & \text{if } x > 0, y > 0 \\ 0 & \text{otherwise.} \end{cases}$$

Since , it cannot be a joint density function.

Therefore, F is not a joint distribution function.

Definition

- ◆ *Let S be a subset of the plane with area $A(S)$. A point is said to be **randomly selected** from S if for any subset R of S with area $A(R)$, the probability that R contains the point is*
- ◆ This definition is essential in the field of **probability**. By the following examples, we will show how it can help to solve problems readily.

Example 8.7

- ◆ A man invites his fiancée to a fine hotel for a Sunday brunch. They decide to meet in the lobby of the hotel between 11:30 A.M. and 12 noon.
- ◆ If they arrive at random times during this period, what is the probability that they will meet within 10 minutes?

1. Let X and Y be the minutes past 11:30 A.M. that the man and his fiancée arrive at the lobby, respectively.

Let $S = \{(x, y) :$ _____ $\},$ and

$R = \{(x, y) \in S : \quad \quad \quad \}.$

2.
$$P(|X - Y| \leq 10) = \frac{\text{area of } R}{30 \times 30} = \frac{\text{area}(R)}{900}$$

3. From fig 8.2, R is the shaded region

$$\text{area}(R) = \quad \quad \quad = 500$$

$$P(|X - Y| \leq 10) = \frac{500}{900} = \frac{5}{9}$$

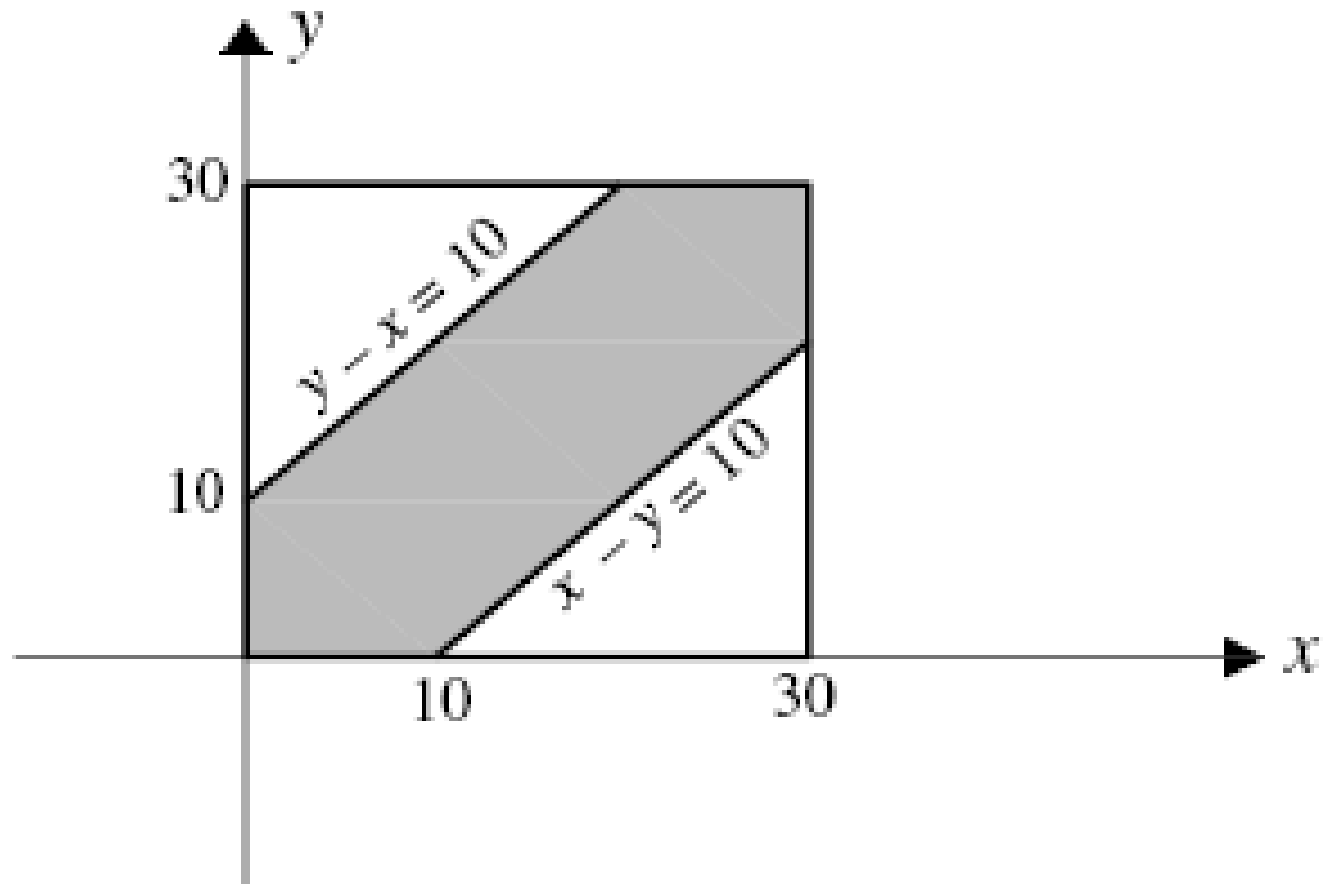


Figure 8.2 Geometric model of Example 8.7.

Theorem 8.2

- ◆ *Let $f(x, y)$ be the joint probability density function of random variables X and Y .*
- ◆ *If h is a function of two variables from \mathbf{R}^2 to \mathbf{R} , then $h(X, Y)$ is a random variable with the expected value given by*

$$E[h(X, Y)] =$$

provided that the integral is absolutely convergent.

Corollary

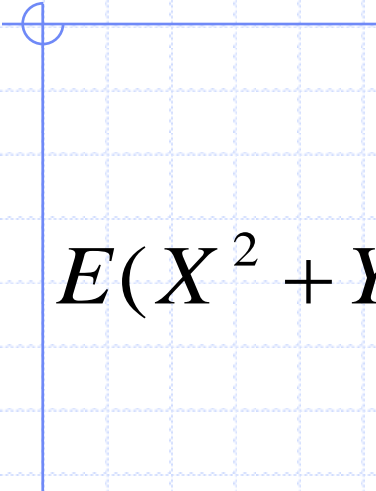
◆ *For random variables X and Y ,*
 $E(X + Y) =$

Example 8.9

- ◆ Let X and Y have joint probability density function

$$f(x, y) = \begin{cases} \frac{3}{2}(x^2 + y^2) & \text{if } 0 < x < 1, 0 < y < 1 \\ 0 & \text{otherwise.} \end{cases}$$

Find $E(X^2 + Y^2)$.


$$E(X^2 + Y^2) =$$

$$= \int_0^1 \int_0^1 \frac{3}{2} (x^2 + y^2)^2 dx dy$$

$$= \frac{3}{2} \int_0^1 \int_0^1 (x^4 + 2x^2 y^2 + y^4) dx dy = \frac{14}{15}.$$

8.2 Independent random variables

- ◆ Two random variables X and Y are called **independent** if, for arbitrary subsets A and B of real numbers, the events $\{X \in A\}$ and $\{Y \in B\}$ are independent, that is, if

$$P(X \in A, Y \in B) =$$

- ◆ This implies that for any two real numbers a and b ,

$$P(X \leq a, Y \leq b) =$$

Theorem 8.3

- ◆ *Let X and Y be two random variables defined on the same sample space.*
- ◆ *If F is the joint probability distribution function of X and Y , then X and Y are independent if and only if for all real numbers t and u ,*

$$F(t,u) =$$

Independence of Discrete Random Variables

◆ Theorem 8.4

- *Let X and Y be two discrete random variables defined on the same sample space.*
- *If $p(x, y)$ is the joint probability mass function of X and Y , then X and Y are independent if and only if for all real numbers x and y ,*

$$p(x, y) = \quad (8.11)$$

◆ Let X and Y be discrete *independent* random variables with sets of possible values A and B , respectively. Then (8.11) implies that for all $x \in A$ and $y \in B$,

$$P(X = x \mid Y = y) =$$

and

$$P(Y = y \mid X = x) =$$

Example 8.10

- ◆ Suppose that 4% of the bicycle fenders, produced by a stamping machine from the strips of steel, need smoothing.
- ◆ What is the probability that, of the next 13 bicycle fenders stamped by this machine, two need smoothing and, of the next 20, three need smoothing?

1. Let X be the number of bicycle fenders among the first 13 that need smoothing.

Let Y be the number of those among the next 7 that need smoothing.

2. We want to calculate $P(X = 2, Y = 1)$. Since X and Y are independent binomial random variables with parameters $(13, 0.04)$ and $(7, 0.04)$, respectively,

$$\begin{aligned} P(X = 2, Y = 1) &= \\ &= \binom{13}{2} (0.04)^2 (0.96)^{11} \binom{7}{1} (0.04)^1 (0.96)^6 \approx 0.0175 \end{aligned}$$

Independence of Continuous Random Variables

◆ Theorem 8.7

- *Let X and Y be jointly continuous random variables with joint probability density function $f(x, y)$.*
- *Then X and Y are independent if and only if $f(x, y)$ is the product of their marginal densities $f_X(x)$ and $f_Y(y)$.*

◆ *By differentiating $F(x, y) = \int_{-\infty}^y \int_{-\infty}^x f(u, v) du dv$, you can prove this theorem!*

Example 8.12

- ◆ Stores A and B , which belong to the same owner, are located in two different towns.
- ◆ If the probability density function of the weekly profit of each store, in thousands of dollars, is given by

$$f(x) = \begin{cases} x/4 & \text{if } 1 < x < 3 \\ 0 & \text{otherwise} \end{cases}$$

and the profit of one store is independent of the other, what is the probability that next week one store makes at least \$500 more than the other store?

1. Let X and Y denote next week's profits of A and B , respectively. The desired probability is

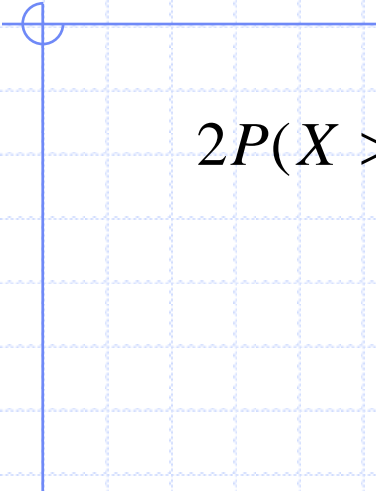
Since X and Y have the same probability density function, by symmetry, this sum equals $2P(X > Y + 1/2)$.

2. $\because X$ and Y are independent

\therefore

$$\because f_X(x) = \begin{cases} x/4 & \text{if } 1 < x < 3 \\ 0 & \text{otherwise} \end{cases} \quad f_Y(y) = \begin{cases} y/4 & \text{if } 1 < y < 3 \\ 0 & \text{otherwise} \end{cases}$$

$$\therefore f(x, y) = \begin{cases} 0 & \text{otherwise} \end{cases}$$


$$2P(X > Y + \frac{1}{2}) =$$

$$= \frac{1}{8} \int_{3/2}^3 \left[\frac{xy^2}{2} \right]_1^{x-1/2} dx$$

$$= \frac{1}{16} \int_{3/2}^3 x \left[\left(x - \frac{1}{2} \right)^2 - 1 \right] dx$$

$$= \frac{1}{16} \int_{3/2}^3 \left(x^3 - x^2 - \frac{3}{4}x \right) dx$$

$$= \frac{1}{16} \left[\frac{1}{4}x^4 - \frac{1}{3}x^3 - \frac{3}{8}x^2 \right]_{3/2}^3 = \frac{549}{1024} \approx 0.54$$

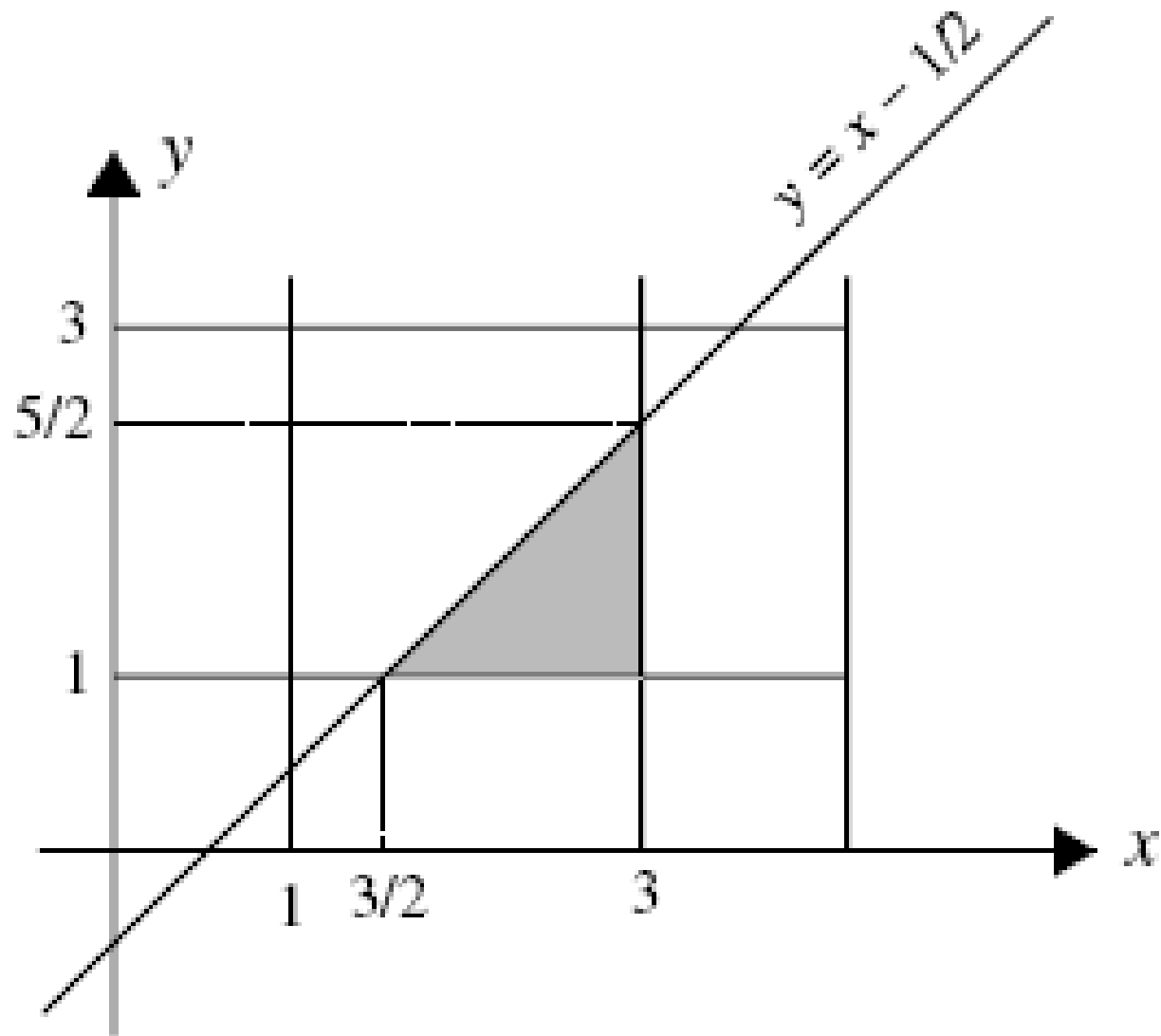


Figure 8.4 Figure of Example 8.12.

Example 8.15

◆ Prove that two random variables X and Y with the following joint probability density function are not independent.

$$f(x, y) = \begin{cases} 8xy & 0 \leq x \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} 1. \quad f_X(x) &= &= 4x(1-x^2), & 0 \leq x \leq 1 \\ f_Y(y) &= &= 4y^3, & 0 \leq y \leq 1 \end{aligned}$$

2. \therefore

$\therefore X$ and Y are dependent

8.3 Conditional distributions

- ◆ Conditional Distributions: Discrete Case
- ◆ Conditional Distributions: Continuous Case

Conditional Distributions: Discrete Case

- ◆ Let X be a discrete random variable with set of possible values A , and let Y be a discrete random variable with set of possible values B .
- ◆ Let $p_{X,Y}(x,y)$ be the joint probability mass function of X and Y , and let $p_X(x)$ and $p_Y(y)$ be the marginal probability mass functions of X and Y .
- ◆ If the value of Y is known, the **conditional probability mass function of X given that $Y = y$** which is denoted by $p_{X|Y}(x|y)$ is defined as follows:

$$p_{X|Y}(x|y) =$$

where $x \in A$, $y \in B$, and $p_Y(y) > 0$.

◆ Note that

$$\sum_{x \in A} p_{X|Y}(x | y) = \frac{1}{p_Y(y)} p_Y(y) = 1$$

◆ Hence for any fixed $y \in B$, $p_{X|Y}(x|y)$ is itself a probability mass function with the set of possible values A .

◆ If X and Y are independent,

$$\begin{aligned} p_{X|Y}(x | y) &= \\ &= P(X = x) = p_X(x) \end{aligned}$$

◆ Similar to $p_{X|Y}(x|y)$, the **conditional distribution function of X , given that $Y = y$** is defined as follows:

$$\begin{aligned} F_{X|Y}(x|y) &= \\ &= \sum_{t \leq x} p_{X|Y}(t|y) \end{aligned}$$

Example 8.16

◆ Let the joint probability mass function of X and Y be given by

$$p(x, y) = \begin{cases} \frac{1}{15}(x + y) & \text{if } x = 0, 1, 2, y = 1, 2 \\ 0 & \text{otherwise} \end{cases}$$

Find $p_{X|Y}(x|y)$ and $P(X = 0 \mid Y = 2)$.

1. $p_y(y) =$

2. $p_{X|Y}(x|y) =$, $x = 0, 1, 2$ when $y = 1$ or 2

$P(X = 0 \mid Y = 2) =$

- ◆ Let X and Y be discrete random variables, and let the set of possible values of X be A .
- ◆ The **conditional expectation of the random variable X given that $Y = y$** is as follows:

$$E(X | Y = y) =$$

where $p_Y(y) > 0$.

◆ If h is a function from \mathbf{R} to \mathbf{R} , then for the discrete random variables X and Y with set of possible values A for X , the expected value of $h(X)$ is obtained from

$$E[h(X) | Y = y] =$$

Example 8.18

◆ Calculate the expected number of aces in a randomly selected poker hand that is found to have exactly two jacks.

1. Let X and Y be the number of aces and jacks in a random poker hand, respectively.
- 2.

$$\begin{aligned} E(X | Y = 2) &= \frac{\sum_{x=0}^3 x \frac{\binom{4}{x} \binom{4}{2-x} \binom{44}{3-x}}{\binom{52}{5}}}{\frac{\binom{4}{2} \binom{4}{3} \binom{44}{3}}{\binom{52}{5}}} = \sum_{x=0}^3 x \frac{\binom{4}{x} \binom{44}{3-x}}{\binom{48}{3}} \approx 0.25 \end{aligned}$$

Conditional Distributions: Continuous Case

- ◆ Let X and Y be two continuous random variables with the joint probability density function $f(x, y)$.
- ◆ When the value of Y is known, to find the probability of events concerning X , $f_{X|Y}(x|y)$, the **conditional probability density function of X given that $Y = y$** is used.
- ◆ $f_{X|Y}(x|y)$ is defined as follows:

$$f_{X|Y}(x, y) =$$

provided that $f_Y(y) > 0$.

◆ Note that

$$\int_{-\infty}^{\infty} f_{X|Y}(x|y)dx = \int_{-\infty}^{\infty} \frac{f(x,y)}{f_Y(y)} dx = \frac{1}{f_Y(y)} f_Y(y) = 1$$

showing that for a fixed y , $f_{X|Y}(x|y)$ is itself a probability density function.

◆ If X and Y are independent, then

$$f_{X|Y}(x|y) = f_X(x)$$

◆ Also, as we expect, $F_{X|Y}(x|y)$, the **conditional probability distribution function of X given that $Y = y$** is defined as follows:

$$F_{X|Y}(x|y) =$$

◆ Therefore,

$$\frac{d}{dx} F_{X|Y}(x|y) =$$

Example 8.21

◆ Let X and Y be continuous random variables with joint probability density function

$$f(x, y) = \begin{cases} \frac{3}{2}(x^2 + y^2) & \text{if } 0 < x < 1, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

Find $f_{X|Y}(x|y)$.

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx = \frac{3}{2} y^2 + \frac{1}{2}$$

$$f_{X|Y}(x | y) = \frac{3(x^2 + y^2)}{3y^2 + 1}, \quad 0 < x < 1 \text{ and } 0 < y < 1$$

Example 8.22

- ◆ First, a point Y is selected at random from the interval $(0, 1)$. Then another point X is chosen at random from the interval $(0, Y)$.
 - ◆ Find the probability density function of X .
1. Let $f(x, y)$ be the joint probability density function of X and Y

$$f_X(x) =$$

$$2. \quad f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} \Leftrightarrow f(x, y) =$$

3. *Y is uniformly distributed over (0, 1),*

$$f_Y(y) = \begin{cases} 1 & \text{if } 0 < y < 1 \\ 0 & \text{elsewhere} \end{cases}$$

4. *given $Y = y$, X is uniformly distributed over $(0, y)$*

$$f_{X|Y}(x|y) = \begin{cases} 1/y & \text{if } 0 < y < 1, \quad 0 < x < y \\ 0 & \text{elsewhere} \end{cases}$$

$$5. \quad f_X(x) = \int_0^1 f_{X|Y}(x|y) dy = \int_x^1 \frac{1}{y} dy = \ln 1 - \ln x = -\ln x$$

$$\therefore f_X(x) = \begin{cases} -\ln x & \text{if } 0 < x < 1 \\ 0 & \text{elsewhere} \end{cases}$$

Example 8.23

◆ Let the conditional probability density function of X , given that $Y = y$, be

$$f_{X|Y}(x|y) = \frac{x+y}{1+y} e^{-x}, \quad 0 < x < \infty, \quad 0 < y < \infty$$

Find $P(X < 1 | Y = 2)$.

1.

$$f_{X|Y}(x|2) = \quad , \quad 0 < x < \infty$$

$$P(X < 1 | Y = 2) =$$

$$= \frac{1}{3} [-xe^{-x} - e^{-x}]_0^1 - \frac{2}{3} [e^{-x}]_0^1 = 1 - \frac{4}{3} e^{-1} \approx 0.509$$

◆ Similar to the case where X and Y are discrete, for continuous random variables X and Y with joint probability density function $f(x, y)$, the **conditional expectation of X given that $Y = y$** is as follows:

$$E(X | Y = y) =$$

where $f_Y(y) > 0$.

◆ If h is a function from \mathbf{R} to \mathbf{R} , then, for continuous random variables X and Y , with joint probability density function $f(x, y)$,

$$E[h(X) | Y = y] =$$

◆ In particular, this implies that the conditional variance of X given that $Y = y$ is given by

$$\sigma_{X|Y=y}^2 =$$

Example 8.24

◆ Let X and Y be continuous random variables with joint probability density function

$$f(x, y) = \begin{cases} e^{-y} & \text{if } y > 0, 0 < x < 1 \\ 0 & \text{elsewhere} \end{cases}$$

Find $E(X \mid Y = 2)$.

$$1. \quad E[X \mid Y = 2] = \int_0^1 x \frac{f(x, 2)}{f_Y(2)} dx = \int_0^1 x \frac{e^{-2}}{f_Y(2)} dx$$

$$\because f_Y(2) = \int_0^1 e^{-2} dx = e^{-2}$$

$$\therefore E(X \mid Y = 2) = \frac{1}{2}$$

Example 8.25

- ◆ The lifetimes of batteries manufactured by a certain company are identically distributed with probability distribution and probability density functions F and f , respectively.
- ◆ In terms of F , f , and s , find the expected value of the lifetime of an s -hour-old battery.

1. Let X be the lifetime of the s -hour-old battery. We want to calculate $E(X \mid X > s)$. Let $F_{X|X>s}(t) = P(X \leq t \mid X > s)$

2. $E(X \mid X > s) =$

$$\because F_{X|X>s}(t) = \begin{cases} 0 & \text{if } t \leq s \\ \frac{P(s < X \leq t)}{P(X > s)} & \text{if } t > s \end{cases}$$

$$= \begin{cases} 0 & \text{if } t \leq s \\ \frac{F(t) - F(s)}{1 - F(s)} & \text{if } t > s \end{cases}$$

$$\because f_{X|X>s}(t) = \frac{d}{dt} F_{X|X>s}(t) = \begin{cases} 0 & \text{if } t \leq s \\ f(t) & \text{if } t > s \end{cases}$$

$$\Rightarrow E(X \mid X > s) = \frac{1}{1 - F(s)} \int_s^{\infty} t f(t) dt$$